Celebrations for
three influential scholars in Information Theory and Statistics:

- Tom Cover: On the occasion of his 70th birthday
  Coverfest.stanford.edu
  Elements of Information Theory Workshop
  Stanford University, May 16, 2008: TODAY!
- Imre Csiszár: On the occasion of his 70th birthday
  www.renyi.hu/∼infocom
  Information and Communication Conference
  Renyi Institute, Budapest, August 25-28, 2008
- Jorma Rissanen: On the occasion of his 75th birthday
  Festschrift at www.cs.tut.fi/∼tabus/
  presented at the *IEEE Information Theory Workshop*
  Porto, Portugal, May 8, 2008

Principles of Information Theory in Probability and Statistics

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May 16, 2008
Elements of Information Theory Workshop

On the Occasion of the Festival for Tom Cover

Outline of Principles

Outline of Principles

Monotonicity of Information Divergence

- Chain Rule

$$D(P_{X,X'} \| P^*_{X,X'}) = D(P_X \| P^*_X) + E\, D(P_{X'|X} \| P^*_{X'|X})$$
$$= D(P_{X'} \| P^*_{X'}) + E\, D(P_{X|X'} \| P^*_{X|X'})$$

- Markov Chains

$$D(P_{X_n} \| P^*) \leq D(P_{X_m} \| P^*) \quad \text{for } n > m$$

and $\log p_n(X_n)/p^*(X_n)$ is a Cauchy sequence in $L_1(P)$

Monotonicity of Information Divergence

- Nonnegative Martingales ρ_n equal the density of a measure Q_n and can be examined in the same way by the chain rule for $n > m$

$$D(Q_n\|P) = D(Q_m\|P) + \int \left(\rho_n \log \frac{\rho_n}{\rho_m}\right) dP$$

- Thus $D(Q_n\|P)$ is an increasing sequence. When it is bounded ρ_n is a Cauchy sequences in $L_1(P)$ with limit ρ defining a measure Q, also, $\log \rho_n$ is a Cauchy sequence in $L_1(Q)$ and

$$D(Q_n\|P) \nearrow D(Q\|P)$$

Monotonicity of Information Divergence: CLT

- Central Limit Theorem Setting:

  $\{X_i\}$ i.i.d. mean zero, finite variance

  $P_n = P_{Y_n}$ is distribution of $Y_n = \frac{X_1 + X_2 + \ldots X_n}{\sqrt{n}}$

  P^* is the corresponding normal distribution

- Chain Rule: Action is mysterious in this case

$$D(P_{Y_m,Y_n} \| P^*_{Y_m,Y_n}) = D(P_{Y_m} \| P^*) + ED(P_{Y_n|Y_m} \| P^*_{Y_n|Y_m})$$

$$= D(P_{Y_n} \| P^*) + ED(P_{Y_m|Y_n} \| P^*_{Y_m|Y_n})$$

Monotonicity of Information Divergence: CLT

- Entropy Power Inequality

$$e^{2H(X+X')} \geq e^{2H(X)} + e^{2H(X')}$$

yields

$$D(P_{2n}\|P^*) \leq D(P_n\|P^*)$$

- Information Theoretic proof of CLT (B. 1986):

$$D(P_n\|P^*) \to 0 \text{ iff finite}$$

Monotonicity of Information Divergence: CLT

- Entropy Power Inequality

$$e^{2H(X+X')} \geq e^{2H(X)} + e^{2H(X')}$$

- Generalized Entropy Power Inequality (Madiman&B.2006)

$$e^{H(X_1+\ldots+X_n)} \geq \frac{1}{r(\mathcal{S})} \sum_{s\in\mathcal{S}} e^{2H(\sum_{i\in s} X_i)}$$

- Proof: simple L_2 projection properties of entropy derivative.
- Consequence, for all $n > m$,

$$D(P_n \| P^*) \leq D(P_m \| P^*)$$

B. and Madiman 2006, Tolino and Verdú 2006
Earlier elaborate proof by Artstein, Ball, Barthe, Naor 2004.

Outline of Principles

AEP and behavior of optimal tests

- Stability of log-likelihood ratios,

$$\frac{1}{n}\log\frac{p(Y_1, Y_2, \ldots Y_n)}{q(Y_1, Y_2, \ldots, Y_n)} \to \mathcal{D}(P\|Q) \text{ with } P - prob\ 1$$

  where $\mathcal{D}(P\|Q)$ is the relative entropy (I-divergence) rate.
- Implication: Associated log-likelihood ratio test A_n has asymptotic P-power 1 (at most finitely many mistakes $P(A_n^c\ i.o.) = 0$) and has optimal Q-prob of error

$$Q(A_n) = \exp\{-n[\mathcal{D} + o(1)]\}$$

- Most general known form of the Chernoff-Stein Lemma.
- Intrinsic role for information divergence rate

$$\mathcal{D}(P\|Q) = \lim \frac{1}{n} D(P_{\underline{Y}^n}\|Q_{\underline{Y}^n})$$

Large Deviations for Empirical Prob and Conditional Limit

- P^*: Information Projection of Q onto convex C
- Pythagorean Identity (Csiszar 75, Topsoe 79): For P in C

$$D(P\|Q) \geq D(C\|Q) + D(P\|P^*)$$

where

$$D(C\|Q) = \inf_{P \in C} D(P\|Q)$$

- Empirical Distribution P_n
- If $D(\text{interior}C\|Q) = D(C\|Q)$ then

$$Q\{P_n \in C\} = \exp\{-n[D(C\|Q) + o(1)]\}$$

and the conditional distribution $P_{Y_1,Y_2,\ldots,Y_n|\{P_n \in C\}}$ converges to $P^*_{Y_1,Y_2,\ldots,Y_n}$ in the I-divergence rate sense (Csiszar 1985)

Outline of Principles

Information Capacity

- A Channel $\theta \longrightarrow \underline{Y}$ is a family of probability distributions

$$\{P_{\underline{Y}|\theta} : \theta \in \Theta\}$$

- Information Capacity

$$C = \max_{P_\theta} I(\theta; \underline{Y})$$

Communications Capacity

- C_{com} = maximum rate of reliable communication, measured in message bits per use of a channel

- Shannon Channel Capacity Theorem (Shannon 1948)

$$C_{com} = C$$

Data Compression Capacity

- Minimax Redundancy

$$Red = \min_{Q_{\underline{Y}}} \max_{\theta \in \Theta} D(P_{\underline{Y}|\theta} \| Q_{\underline{Y}})$$

- Data Compression Capacity Theorem

$$Red = C$$

(Gallager, Davisson & Leon-Garcia, Ryabko)

Statistical Risk Setting

- Loss function

$$\ell(\theta, \theta')$$

- Examples:
- Kullback Loss

$$\ell(\theta, \theta') = D(P_{Y|\theta} \| P_{Y|\theta'})$$

- Squared metric loss, e.g. squared Hellinger loss:

$$\ell(\theta, \theta') = d^2(\theta, \theta')$$

Statistical Capacity

- Estimators: $\hat{\theta}_n$
- Based on sample $\underline{Y}$ of size n
- Minimax Risk (Wald):

$$r_n = \min_{\hat{\theta}_n} \max_{\theta} E\ell(\theta, \hat{\theta}_n)$$

Ingredients for determination of Statistical Capacity

- Kolmogorov Metric Entropy of $S \subset \Theta$:

$$H(\epsilon) = \max\{\log \mathit{Card}(\Theta_\epsilon) \ : \ d(\theta, \theta') > \epsilon \text{ for } \theta, \theta' \in \Theta_\epsilon \subset S\}$$

- Loss Assumption, for $\theta, \theta' \in S$:

$$\ell(\theta, \theta') \sim D(P_{Y|\theta} \| P_{Y|\theta'}) \sim d^2(\theta, \theta')$$

Statistical Capacity Theorem

- For infinite-dimensional Θ
- With metric entropy evaluated a critical separation ϵ_n
- Statistical Capacity Theorem

  Minimax Risk $\sim$ Info Capacity Rate $\sim$ Metric Entropy rate

$$r_n \quad \sim \quad \frac{C_n}{n} \quad \sim \quad \frac{H(\epsilon_n)}{n} \quad \sim \quad \epsilon_n^2$$

Yang 1997, Yang and B. 1999, Haussler and Opper 1997

Outline of Principles

Shannon Codes

- Kraft-McMillan characterization:
  Uniquely decodeable codelengths

$$L(\underline{x}), \quad \underline{x} \in \underline{\mathcal{X}}, \qquad \sum_{\underline{x}} 2^{-L(\underline{x})} \leq 1$$

$$L(\underline{x}) = \log 1/p(\underline{x}) \qquad p(\underline{x}) = 2^{-L(\underline{x})}$$

- Operational meaning of probability:

  A probability distribution p is given by a choice of code

Complexity

- Kolmogorov Idealized Compression:

$$K(\underline{Y}) = \text{Length of shortest computer code for } \underline{Y} \text{ on a given universal computer}$$

- Shannon Idealized Codelength (expectation optimal):

$$\log 1/p^*(\underline{Y})$$

- But p^* is not generally known
- Hybrid: Statistical measure of Complexity of $\underline{Y}$

$$\min_p \left[\log 1/p(\underline{Y}) \; + \; L(p) \right]$$

- Valid for any $L(p)$ satisfying Kraft summability.

Complexity

- Minimum Description Length Principle (MDL)
- Special cases: Rissanen 1978,1983,...
- Two-stage codelength formulation:
  (Cover Scratch pad 1981, B. 1985, B. and Cover 1991)

$$L(\underline{Y}) = \min_p \Big[\quad \log 1/p(\underline{Y}) \quad + \quad L(p) \quad \Big]$$

bits for $\underline{x}$ given p + bits for p

- Corresponding statistical estimator $\hat{p}$ achieves the above minimization in a family of distributions for a specified $L(p)$

MDL Analysis

- Redundancy of Two-stage Code:

$$Red = \frac{1}{n} E \left\{ \min_p \left[\log \frac{1}{p(\underline{Y})} + L(p) \right] - \log \frac{1}{p^*(\underline{Y})} \right\}$$

- bounded by Index of Resolvability:

$$Res_n(p^*) = \min_p \left\{ D(p^* || p) + \frac{L(p)}{n} \right\}$$

- Statistical Risk Analysis in i.i.d. case:

$$E\, d^2(p^*, \hat{p}) \leq \min_p \left\{ D(p^* \| p) + \frac{L(p)}{n} \right\}$$

- B. 1985, B.&Cover 1991, B., Rissanen, Yu 1998, Li 1999, Grunwald 2007

MDL Analysis

- Statistical Risk Analysis:

$$E\, d^2(p^*, \hat{p}) \leq \min_p \{D(p^* \| p) \; + \; L(p)/n\}$$

- Special Cases:
  Traditional parametric: $L(\theta) = (dim/2) \log n + C$
  Nonparametric: $L(p) =$ Metric entropy
  (log cardinality of optimal net)
  Idealized: $L(p) =$ Kolmogorov complexity
- Adaptation:
  Achieves minimax optimal rates simultaneously in every computable subfamily of distributions

MDL Analysis: Key to risk consideration

- Discrepancy between training sample and future

$$Disc(p) = \log \frac{p^*(\underline{Y})}{p(\underline{Y})} - \log \frac{p^*(\underline{Y}')}{p(\underline{Y}')}$$

- Future term may be replaced by population counterpart
- Discrepancy control: If $L(p)$ satisfies the Kraft sum then

$$E\left[\sup_p \{Disc(p) + 2L(p)\}\right] \geq 0$$

- From which the risk bound follows:

$$\text{Risk} \leq \text{Redundancy} \leq \text{Resolvability}$$

$$E\, d^2(p^*, \hat{p}) \leq Red \leq Res_n(p^*)$$

Statistically valid penalized likelihood

- New result
- (B., Li, Huang, Luo 2008, Festschrift for Jorma Rissanen)
- Penalized Likelihood:

$$\hat{p} = \arg\min_p \left\{ \frac{1}{n} \log \frac{1}{p(Y)} \; + \; pen_n(p) \right\}$$

- Penalty condition:

$$pen_n(p) \geq \frac{1}{n} \min_{\tilde{p}} \{ 2L(\tilde{p}) \; + \; \Delta_n(p, \tilde{p}) \}$$

where the distortion $\Delta_n(p, \tilde{p})$ is the difference in discrepancies at p and a representer $\tilde{p}$

- Risk conclusion:

$$Ed^2(p, \hat{p}) \leq \inf_p \{ D(p^* \| p) \; + \; pen_n(p) \}$$

Example: ℓ_1 penalties on coefficients in gen. linear models

- $\mathcal{G}$ is a dictionary of candidate basis functions
- Wavelets, splines, polynomials, trigonometric terms, sigmoids, explanatory variables and their interactions
- Candidate functions in the linear span

$$f(x) = f_\theta(x) = \sum_{g \in \mathcal{G}} \theta_g \, g(x)$$

- ℓ_1 norm of coefficients

$$\|\theta\|_1 = \sum_g |\theta_g|$$

Example Models

- $p = p_f$ specified through a family of function f
- Regression

$$p_f(y|x) = \text{Normal}(f(x), \sigma^2)$$

- Logistic regression with $y \in \{0, 1\}$

$$p_f(y|x) = \text{Logistic}(f(x)) \quad \text{for } y = 1$$

- Log-density estimation

$$p_f(x) = \frac{p_0(x) \exp\{f(x)\}}{c_f}$$

ℓ_1 Penalty

- $pen_n(f_\theta) = \lambda_n \|\theta\|_1$ where $f_\theta(x) = \sum_{g \in \mathcal{G}} \theta_g \, g(x)$
- Popular penalty: Chen & Donoho (96) Basis Pursuit; Tibshirani (96) LASSO; Efron et al (04) LARS; Precursors: Jones (92), B.(90,93,94) greedy algorithm and analysis of combined ℓ_1 and ℓ_0 penalty
- Risk analysis: specify valid λ_n for risk $\leq$ resolvability
- Computation analysis: bound accuracy of new ℓ_1-penalized greedy pursuit algorithm

ℓ_1 penalty is valid for λ_n of order $1/\sqrt{n}$

- Example: ℓ_1 penalized log-density estimation, i.i.d. case

$$\hat{\theta} = \text{argmin}_\theta \left\{ \frac{1}{n} \log \frac{1}{p_{f_\theta}(\underline{x})} + \lambda_n \|\theta\|_1 \right\}$$

- Risk bound:

$$Ed(f^*, f_{\hat{\theta}}) \leq \inf_\theta \left\{ D(f^* || f_\theta) + \lambda_n \|\theta\|_1 \right\}$$

- Valid for $\lambda_n \geq B\sqrt{\frac{H}{n}}$ with $H = \log Card(\mathcal{G})$
- $B =$ bound on the range of functions in $\mathcal{G}$
- For infinite cardinality $\mathcal{G}$ use metric entropy in place of H
- Results for regression shown in a companion paper

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- Risk bound:

$$Ed(f^*, f_{\hat{\theta}}) \leq \inf_\theta \left\{ D(f^* || f_\theta) + \lambda_n \|\theta\|_1 \right\}$$

- True with

$$\lambda_n = B\sqrt{\frac{H}{n}} \quad \text{with } H = \log Card(\mathcal{G})$$

- Risk of order λ_n when the target has finite ℓ_1 norm

Comment on proof

- Shannon-like demonstration of the existence of the variable complexity cover property
- Inspiration from technique originating with Lee Jones (92)
- Representer $\tilde{f}$ of f_θ of the form

$$\tilde{f}(x) = \frac{v}{m} \sum_{k=1}^{m} g_k(x)$$

- $g_1, \ldots g_m$ picked at random from $\mathcal{G}$, independently, where g arises with probability proportional to $|\theta_g|$
- May pick them in greedy fashion as in demonstration of fast computation properties
- In the paper in the Festschrift for Rissanen with summary in the ITW - Porto proceedings

Outline of Principles

Summary

- Information divergence is monotone
- Info Capacity $\sim$ Minimax Redundancy $\sim$ Minimax Risk
- Adaptivity of MDL: simultaneously minimax optimal
- Penalized Likelihood risk analysis
- ℓ_0 penalty $\frac{dim}{2} \log n$ classically analyzed
- ℓ_1 penalty $\lambda_n \|\theta\|_1$ valid for $\lambda_n \geq \sqrt{H/n}$.